



Original Research

Hypertensive Disorders of Pregnancy and Maternal Long-Term Kidney Dysfunction: A Mendelian Randomization Analysis

Huanqiang Zhao^{1,†}, Shuyi Shao^{1,†}, Yang Zi¹, Ping Wen¹, Qixin Xu¹, Xiujie Zheng¹, Shiguo Chen¹, Zhiheng Wang^{2,3}, Yu Xiong^{2,3,*}, Xiaotian Li^{1,*}

¹Shenzhen Maternity and Child Healthcare Hospital, Women and Children's Medical Center, Southern Medical University, 518028 Shenzhen, Guangdong, China

²Obstetrics and Gynecology Hospital, Fudan University, 200011 Shanghai, China

³The Shanghai Key Laboratory of Female Reproductive Endocrine-Related Diseases, 200011 Shanghai, China

*Correspondence: xiongyu1535@163.com (Yu Xiong); xli555@fudan.edu.cn (Xiaotian Li)

†These authors contributed equally.

Academic Editors: George Daskalakis and Michael H. Dahan

Submitted: 12 July 2025 Revised: 28 September 2025 Accepted: 14 October 2025 Published: 16 January 2026

Abstract

Background: Observational studies have demonstrated a potential association between hypertensive disorders of pregnancy (HDPs) and an elevated risk of subsequent kidney impairment. The present study aimed to evaluate the possible causal relationship between HDPs and future renal dysfunction using a genetic approach. **Methods:** This two-sample Mendelian randomization (MR) analysis, conducted between October 2023 and January 2024, examined the effects of HDPs—both overall and by subtype (gestational hypertension and preeclampsia/eclampsia)—on chronic kidney disease (CKD), albuminuria, and estimated glomerular filtration rate (eGFR). Genetic data were sourced from publicly available genome-wide association studies (GWAS) provided by the FinnGen, CKDGen, UK Biobank, GIANT (Genetic Investigation of ANthropometric Traits), deCODE, and DIAMANTE (DIAbetes Meta-ANalysis of Trans-Ethnic association studies) consortia. The primary method employed was inverse-variance weighting, supplemented by sensitivity analyses and instrument-strength validation to ensure causal robustness. **Results:** This study included approximately 300,000 to 600,000 individuals per kidney function trait, with around 15,000 HDP cases. Genetically predicted HDPs showed a weak but statistically significant association with albuminuria [odds ratio (OR): 1.02; 95% confidence interval (CI): 1.00–1.03; $p = 0.01$]. This association remained consistent after adjustments for body mass index (BMI), smoking, and type 2 diabetes (T2D) in the multivariable MR analysis (OR: 1.03; 95% CI: 1.01–1.04; $p < 0.001$). No significant associations were observed between HDPs, including for the subtypes, and the presence of CKD or eGFR. These results demonstrated robustness across diverse sensitivity analyses. **Conclusions:** The genetic proxy for HDPs was identified as causally associated with maternal long-term albuminuria, independent of BMI, smoking, and T2D. These results offer novel insights bolstering the causal relationship between HDPs and long-term maternal kidney dysfunction.

Keywords: hypertensive disorders of pregnancy; preeclampsia; long-term kidney function; albuminuria; Mendelian randomization

1. Introduction

Hypertensive disorders of pregnancy (HDPs) comprising conditions such as preeclampsia and gestational hypertension affect approximately 10% of pregnancies worldwide. It is associated with substantial immediate maternal morbidity and mortality [1], as well as susceptibility to long-term chronic diseases, such as chronic kidney damage, which has been a common and well-documented concern [2,3].

Acute renal damage represents a frequent manifestation of HDPs and often resolves within months postpartum. Nevertheless, various population-based observational studies have established a correlation between HDPs and long-term kidney diseases [4–8]. A secondary analysis of data from the Family Blood Pressure Program indicated that a history of hypertension during pregnancy was associated with an increased risk of microalbuminuria decades later [9]. Furthermore, analyses of data from the Medi-

cal Birth Registry of Norway and Swedish Medical Birth Register (MBR) demonstrated a link between preeclampsia and an increased risk of end-stage renal disease (ESRD), although the absolute risk remains low [10,11]. In addition, meta-analysis of 18 observational studies revealed a significant positive correlation between two subtypes of HDPs, namely preeclampsia and gestational hypertension, and an increased risk of future chronic kidney disease (CKD) and ESRD [6]. However, it remains unclear whether these associations express causal relationships deriving from the adverse effects of preeclampsia itself, or if they can be attributed to underlying risk factors that predisposed women to both preeclampsia and subsequent renal diseases.

Mendelian randomization (MR) utilizes genetic variants as instrumental variables to evaluate potential causal effects of exposures on outcomes. This method leverages the inherent causal relationship between genetic variants and the exposure of interest. Because genetic alleles are



Copyright: © 2026 The Author(s). Published by IMR Press.
This is an open access article under the CC BY 4.0 license.

Publisher's Note: IMR Press stays neutral with regard to jurisdictional claims in published maps and institutional affiliations.

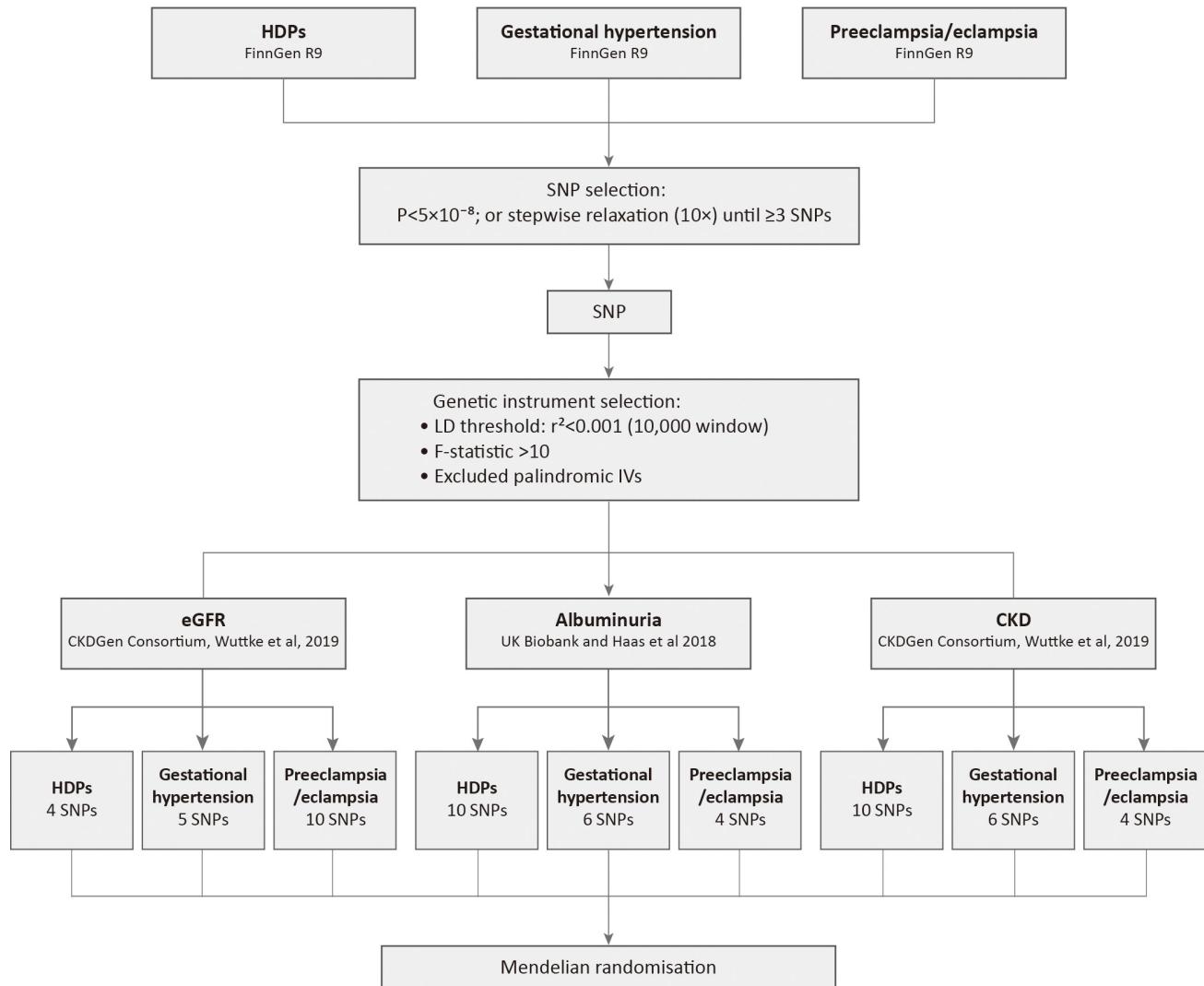


Fig. 1. Study analysis flowchart: from data to instrumental variants. HDPs, hypertensive disorders of pregnancy; CKD, chronic kidney disease; SNP, single nucleotide polymorphism; LD, linkage disequilibrium; eGFR, estimated glomerular filtration rate; IVs, instrumental variables.

randomly assigned at conception and remain largely unaffected by confounding factors that commonly plague observational studies, MR simulates the design of a randomized controlled trial. This unique feature effectively minimizes confounding and reverse causation [12]. Consequently, MR yields estimates that represent the lifelong effect of genetic predisposition to an exposure—such as HDPs—on outcomes like kidney impairment. It is essential to note that these effect estimates are not directly comparable to hazard ratios from conventional observational studies, which generally assess short-term associations over limited follow-up durations and are susceptible to unmeasured confounding. Nonetheless, MR’s capacity to approximate a natural randomized experiment underscores its strength in reinforcing causal inference.

Using a two-sample MR framework, this study examined the causal effects of HDPs, overall and by subtype, on CKD, albuminuria, and reduced estimated glomerular fil-

tration rate (eGFR). To control for confounding, multivariable MR included adjustments for body mass index (BMI), smoking, and type 2 diabetes (T2D).

2. Materials and Methods

2.1 Study Design and Data Sources

This study utilized de-identified, publicly available genome-wide association studies (GWAS) summary data. All original studies had obtained ethical approval. The reporting of this study adheres to the Strengthening the Reporting of Observational Studies in Epidemiology Using Mendelian Randomization (STROBE-MR) guidelines. Data were sourced from studies that individually analyzed populations of European ancestry to mitigate bias stemming from demographic stratification.

Fig. 1 illustrates the overall study design in a flowchart. This analysis focused on three primary expo-

sures: HDPs overall, along with two subtypes—gestational hypertension and preeclampsia/eclampsia. Summary-level genetic association data for these phenotypes were sourced from the FinnGen (<https://www.finngen.fi/en>) consortium's ninth release, which was published in late 2023 [13]. The sample sizes included 14,727 cases and 196,143 controls for any HDPs, 8502 cases and 194,266 controls for gestational hypertension, and 7212 cases and 194,266 controls for preeclampsia/eclampsia. The case definitions for HDPs, gestational hypertension, and preeclampsia/eclampsia were determined using specific phenotype codes within the FinnGen consortium: O15_HYPHENSPREG, O15_GESTAT_HYPERT, and O15_PRE_OR_ECLAMPSIA, respectively. These codes are derived from the International Classification of Diseases (ICD) system, providing a standardized and reliable framework for phenotype classification, thereby enhancing the validity of the analysis (Table 1, Ref. [13–18]).

The MR design estimates the lifetime risk of an outcome associated with genetic predisposition to an exposure. The kidney outcome data (eGFR, albuminuria, CKD) were obtained from GWAS involving adult populations, with measurements taken at a single time point during study enrollment (Table 1). Summary statistics for eGFR (n = 567,460 individuals) and CKD (n = 480,698 individuals, 41,395 cases and 439,303 controls) were derived from meta-analysis of GWAS of participants of European ancestry from the CKDGen Consortium [14]. Participant characteristics and the phenotype definitions have been comprehensively described by Wutke *et al.* [14]. CKD was specifically defined as an eGFR <60 mL/min/1.73 m². eGFR was estimated using the Chronic Kidney Disease Epidemiology Collaboration (CKD-EPI) equation for individuals above 18 years of age [19] and the Schwartz formula for those aged 18 years or younger [20]. The GWAS summary statistics for albuminuria in individuals of European ancestry were obtained from a cohort of 382,500 participants from the UK Biobank [15]. Albuminuria levels were measured during the initial enrollment visit using a Beckman Coulter AU5400 clinical chemistry analyzer with a detection range of 6.7–200 mg/L [15].

Three variables—BMI, smoking, and T2D—were selected as potential confounders when evaluating the association between HDPs and subsequent kidney impairment (Table 1). These factors were included due to their well-documented, although not necessarily causal, associations with both the exposures and outcomes of interest [21–23]. Genetic instruments for each confounder were derived from GWAS with the largest available sample sizes, focusing on individuals of European ancestry. Specifically, summary-level data for BMI were acquired from a meta-analysis including 434,794 female participants, conducted jointly by the UK Biobank and the GIANT (Genetic Investigation of ANthropometric Traits) consortium [16]. Genetic associations for T2D were obtained from a trans-ethnic

GWAS meta-analysis by the DIAMANTE (DIAbetes Meta-ANalysis of Trans-Ethnic association studies) Consortium, comprising 80,154 cases and 853,816 controls [17]. Data for smoking initiation, based on 557,337 cases and 674,754 controls, were collected from the deCODE study and the UK Biobank [18].

As this study utilized summary-level GWAS data, individual-level characteristics and inclusion/exclusion criteria were determined by the original consortia. Details can be found in the referenced publications.

2.2 Instrumental Variants for Exposures

Genetic instruments were selected if associated with each exposure at a genome-wide significance threshold of $p < 5 \times 10^{-8}$. However, because fewer than three independent single-nucleotide variants (SNVs) were identified for gestational hypertension and for preeclampsia/eclampsia at this threshold, the significance criterion was relaxed incrementally by a factor of 10 until at least three SNVs could be included. Consequently, the following adjusted thresholds were applied: $p < 5 \times 10^{-7}$ for gestational hypertension in relation to eGFR; $p < 5 \times 10^{-6}$ for preeclampsia/eclampsia in relation to eGFR; $p < 5 \times 10^{-7}$ for preeclampsia/eclampsia in relation to albuminuria; and $p < 5 \times 10^{-7}$ for preeclampsia/eclampsia in relation to CKD. Genetic instruments for the MR analysis are listed in **Supplementary Table 1**.

The selected genetic instruments were grouped for independence by estimating linkage disequilibrium (LD), applying a cut-off of $r^2 < 0.001$ and a window size of 10,000 kb. In cases where single nucleotide polymorphisms (SNPs) were in LD, we retained those with the lowest p value.

2.3 Testing Instrument Strength and Statistical Power

The strength of the chosen genetic instruments was assessed through the calculation of F statistics, and genetic instruments with an F statistic below 10 were deemed weak and therefore excluded. The evaluation was based on the mathematical formula utilized: $F = R^2 \times (N - K - 1) / [K \times (1 - R^2)]$, where R^2 represents the variance in exposure, N stands for the sample size, and K represents the number of SNPs used in the MR analysis.

To determine the statistical power of our primary univariable MR analyses, we utilized the online tool mRnd (<https://shiny.cnsgenomics.com/mRnd/>), which is specifically designed for power calculations in MR studies [24,25]. This tool estimates statistical power using key parameters such as sample size, case-control ratio, and the proportion of variance in the exposure explained by the genetic instruments.

2.4 MR Analysis

The inverse variance weighting (IVW) method was employed as the primary approach to infer potential causal

Table 1. Overview of genome-wide association studies (GWAS) used.

Trait	Study/Source	Participants	Phenotype definition
Exposure			
HDPs		Cases: 14,727; Control: 196,143	Cases: FinnGen code O15_HYPTENSPREG Controls: not classified as a case
Gestational hypertension	FinnGen R9 [13]	Cases: 8502; Control: 194,266	Cases: FinnGen code O15_GESTAT_HYPERT Controls: not classified as a case
Preeclampsia/eclampsia		Cases: 7212; Control: 194,266	Cases: FinnGen code O15_PRE_OR_ECLAMPSIA Controls: not classified as a case
Outcome			
eGFR	CKDGen Consortium and Wuttke <i>et al.</i> , 2019 [14]	567,460 individuals	eGFR was estimated using the CKD-EPI equation for individuals above 18 years of age and the Schwartz formula for those aged 18 years or younger.
Albuminuria	UK Biobank and Haas <i>et al.</i> , 2018 [15]	382,500 individuals	Albuminuria levels were measured during the initial enrollment visit using a Beckman Coulter AU5400 clinical chemistry analyzer with a detection range of 6.7–200 mg/L.
CKD	CKDGen Consortium and Wuttke <i>et al.</i> , 2019 [14]	Cases: 41,395; Control: 439,303	Cases: eGFR below 60 mL/min/1.73 m ² Controls: not classified as a case
Covariate			
BMI	UK Biobank, GIANT consortium and Pulit <i>et al.</i> , 2019 [16]	434,794 female individuals	Body mass index
Smoking	deCODE, UK Biobank, and Liu <i>et al.</i> , 2019 [18]	Cases: 557,337; Controls: 674,754	Cases: self-reported smoking behavior Controls: never smoked
T2D	DIAMANTE Consortium and Mahajan <i>et al.</i> , 2022 [17]	Cases: 80,154; Controls: 853,816	Cases: T2D was defined based on the criteria outlined in the previous study Controls: not classified as a case

BMI, body mass index; T2D, type 2 diabetes; GIANT, Genetic Investigation of ANthropometric Traits; DIAMANTE, DIAbetes Meta-ANalysis of Trans-Ethnic association studies; CKD-EPI, Chronic Kidney Disease Epidemiology Collaboration.

relationships between HDPs and kidney damage. IVW is known for its ability to offer robust causal estimates free from directional pleiotropy, making it a commonly used method in MR studies [26]. The combined causal estimate from the IVW method represents a weighted average of the individual SNP effects, which may naturally vary in direction due to the complex genetic architecture of the exposure.

2.5 Sensitivity Analyses

To assess the reliability of the main causal estimates and identify potential pleiotropic effects, we implemented several sensitivity analyses, including weighted median MR, MR-Egger regression, and MR-PRESSO (Pleiotropy RESidual Sum and Outlier). The weighted median approach produces consistent causal estimates, provided that at least half of the IVs are valid [27]. MR-Egger regression enables the detection of directional pleiotropy via its intercept term and yields pleiotropy-robust effect estimates [28]. MR-PRESSO detects outlier SNPs and provides effect estimates adjusted from these outliers [29]. Heterogeneity across IVs was assessed using Cochran's Q statistic [30]. A significance level of $p > 0.05$ indicated the absence of noteworthy heterogeneity or pleiotropy among the scrutinized SNPs.

2.6 Multivariable MR Analyses

Multivariable MR utilized SNPs from GWAS on BMI, smoking, and T2D, applying IVW to estimate their concurrent effects. This method re-evaluates summary genetic associations by incorporating genetic correlations with both exposure and risk factors into a weighted regression framework, thereby reducing the influence of SNP exposure on their effects on other presumed risk factors through an indirect pathway. All statistical analyses were performed in R (version 4.3.1, R Foundation for Statistical Computing, Vienna, Austria) with the TwoSampleMR, MR-PRESSO, and MVMR (Multivariable Mendelian Randomization) packages.

3. Results

3.1 MR Analyses

The complete set of genetic instruments employed for the exposures is detailed in **Supplementary Table 1**. The F-statistics for these genetic instruments exceeded the commonly selected threshold of 10, indicating robust instrument strength.

3.2 Associations of HDPs With the Risk of CKD

Contrary to conventional analyses that suggest a positive association between HDPs and the risk of CKD, the MR analysis did not provide evidence to support such associations between HDPs (OR: 1.01; 95% CI: 0.93–1.09; $p = 0.78$) and the subtypes (gestational hypertension: OR: 1.02; 95% CI: 0.94–1.11; $p = 0.63$; preeclampsia/eclampsia: OR: 0.97; 95% CI: 0.81–1.14; $p = 0.68$; **Supplementary Table 2** and Fig. 2) with the risks of CKD.

MR-Egger regression analysis indicated an absence of directional pleiotropy for genetically predicted HDPs (intercept $p = 0.92$; **Supplementary Table 3**). Consistent results were observed for both subtypes: gestational hypertension (intercept $p = 0.42$) and preeclampsia/eclampsia (intercept $p = 0.34$). Effect estimates derived from the weighted median method were consistent with those obtained from the primary IVW analysis, supporting the robustness of the findings. Some heterogeneity was detected for preeclampsia/eclampsia (Cochran's Q = 8.88, $p = 0.03$; **Supplementary Table 3**). In contrast, neither HDPs overall (Q = 5.17, $p = 0.64$) nor gestational hypertension (Q = 2.34, $p = 0.67$) exhibited significant heterogeneity. Furthermore, MR-PRESSO analysis revealed no evidence that outliers substantially influenced the causal estimates (**Supplementary Table 4**).

3.3 Associations of HDPs With Albuminuria

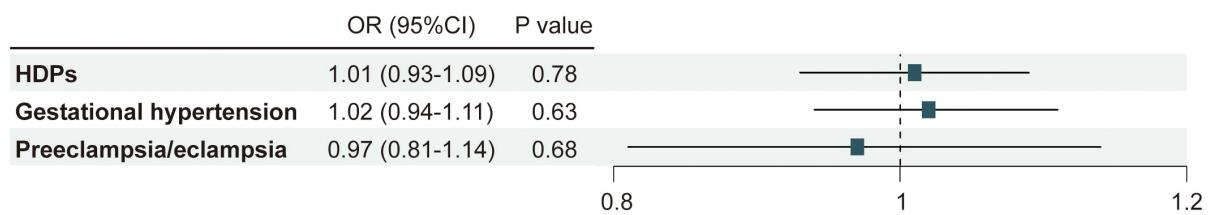
Genetically predicted HDPs demonstrated a significant yet weak association with an elevated risk of albuminuria (OR: 1.02; 95% CI: 1.00–1.03; $p = 0.01$). In contrast, neither gestational hypertension (OR: 1.02; 95% CI: 0.998–1.03; $p = 0.08$) nor preeclampsia/eclampsia (OR: 1.03; 95% CI: 0.99–1.06; $p = 0.13$) showed statistically significant associations with albuminuria, as detailed in **Supplementary Table 2** and Fig. 2.

As shown in **Supplementary Table 3**, some heterogeneity was noted for preeclampsia/eclampsia (Q statistic = 10.63; $p = 0.01$), however, not for HDPs (Q statistic = 7.85; $p = 0.35$) or gestational hypertension (Q statistic = 5.85; $p = 0.21$). MR-PRESSO analysis identified no outlier SNPs (**Supplementary Table 4**), indicating that the heterogeneity likely reflects diffuse mechanisms rather than a few invalid instruments. The MR-Egger test did not detect significant pleiotropy (MR-Egger intercept $p = 0.42$, 0.28, and 0.11 for HDPs, gestational hypertension, and preeclampsia/eclampsia, respectively). The weighted median analysis indicated a consistent direction of effect between genetically predicted HDPs and albuminuria, aligning with the results obtained from the IVW method.

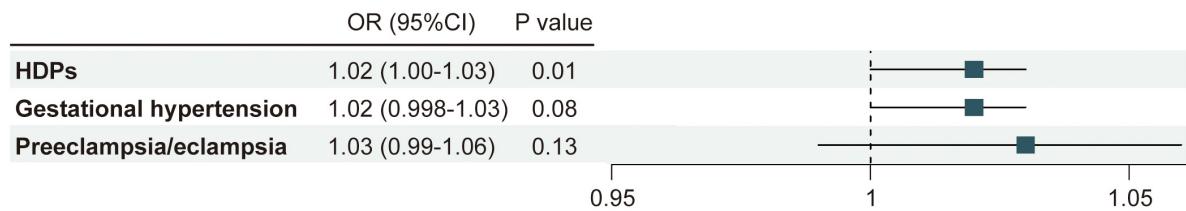
3.4 Associations of HDPs With eGFR

No significant associations were observed between genetically predicted HDPs and eGFR, as indicated by the following ORs: HDPs overall (OR: 1.00; 95% CI: 0.99–1.02; $p = 0.61$), gestational hypertension (OR: 1.00; 95% CI: 0.98–1.02; $p = 0.89$), and preeclampsia/eclampsia (OR: 1.01; 95% CI: 0.999–1.01; $p = 0.10$) (**Supplementary Table 2**; Fig. 2). Sensitivity analyses provided no evidence of directional pleiotropy for HDPs (MR-Egger intercept $p = 0.48$) or its subtypes (gestational hypertension: $p = 0.16$; preeclampsia/eclampsia: $p = 0.55$) (**Supplementary Tables 3,4**). Moreover, the weighted median estimates showed close agreement with the IVW results in both direction and magnitude of effect. Some heterogeneity was

A Outcome: chronic kidney disease



B Outcome: albuminuria



C Outcome: estimated glomerular filtration rate

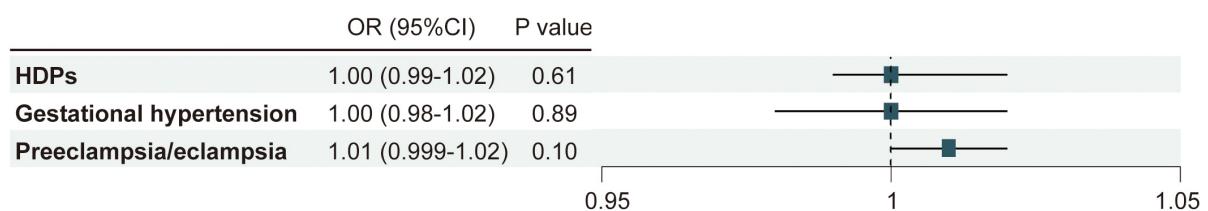


Fig. 2. Associations of genetically predicted HDPs and its subtypes with the risk of kidney function. (A) Associations of HDPs with the risk of CKD. (B) Associations of HDPs with albuminuria. (C) Associations of HDPs with eGFR. OR, odds ratio; CI, confidence interval.

Table 2. Associations of HDPs with albuminuria in multivariable analyses, adjusting for three risk factors.

Exposure	Outcome	Covariate	nSNP	β	Standard error	OR (95% CI)	p value
HDPs	Albuminuria	BMI, smoking, and T2D	6	0.03	0.01	1.03 (1.01-1.04)	<0.001

detected for HDPs (Cochran's $Q = 12.38$; $p = 0.01$) and gestational hypertension ($Q = 12.23$; $p = 0.01$), though not for preeclampsia/eclampsia ($Q = 10.30$; $p = 0.33$). The MR-PRESSO analysis identified 4 outlier SNPs in the association between HDPs and eGFR; however, after their removal, the corrected estimates remained consistent with the initial IVW results, confirming that these outliers did not influence the overall null association.

3.5 Multivariable MR Analysis

Ensuring the accuracy and validity of the results entails addressing potential confounding factors by adjusting variables. Prior epidemiological research has recognized BMI, smoking behavior, and T2D as established confounding variables in the association between HDPs and subsequent renal impairment [21–23]. In the multivariable MR analysis, a consistent and significant association remained

between HDPs and albuminuria (OR: 1.03; 95% CI: 1.01–1.04; $p < 0.001$; Table 2 and **Supplementary Table 5**). Furthermore, the MR-Egger regression analysis revealed no evidence of directional pleiotropy (intercept $p = 0.32$; **Supplementary Table 6**), supporting the robustness of this causal inference.

4. Discussion

4.1 Main Findings

In this MR analysis, we identified a modest yet statistically significant association between the genetic proxy of HDPs and albuminuria, independent of maternal BMI, smoking, and T2D. The results were robust across various sensitivity analyses. Our findings provide novel evidence strengthening the causal association between HDPs and long-term maternal renal dysfunction.

4.2 Comparison and Interpretation

Epidemiological evidence from multiple observational studies indicates that HDPs, particularly preeclampsia, are associated with increased risk of CKD [10,31] and a decline in eGFR in later life [7]. A previous meta-analysis revealed that, on average, 7.1 years postpartum, individuals with severe preeclampsia faced an eight-fold higher risk of microalbuminuria [31]. Another study emphasized that, throughout the postpartum years, a history of HDPs correlated with an elevated risk of microalbuminuria [9]. A recent study has also highlighted a significant long-term escalation in kidney disease risk among women with a history of HDPs [32].

However, establishing causal relationships from observational data presents challenges due to potential residual confounding factors and biases. Many pregnant women with preeclampsia may have undiagnosed preexisting kidney dysfunction, as certain diagnostic tests are not routinely conducted. Additionally, biases might arise from residual confounding in the observational setting, especially given the multifaceted nature of preeclampsia, which is influenced by various sociobehavioral factors [33,34]. This study contributes to existing literature by employing MR, a method that reduces the impact of confounding factors and biases. Findings from this approach revealed a causal relationship between HDPs and long-term maternal albuminuria, a mild manifestation of CKD, although HDPs were not directly associated with CKD or eGFR.

HDPs and CKD may exhibit a bidirectional relationship. Maternal renal disease, even with mild reductions in eGFR or an increase in microalbuminuria, is a recognized risk factor for preeclampsia [35,36]. HDPs, especially preeclampsia, marked by the dysregulation of angiogenic mediators precipitating systemic endothelial dysfunction, can be exacerbated by CKD, altering the complement and renin-angiotensin-aldosterone systems. Our study showed that HDPs, in turn, have been implicated in the onset of long-term albuminuria. HDPs have the potential to instigate future kidney disorders through mechanisms like acute kidney injury, endothelial impairment, and podocyte loss during pregnancy [37,38]. Furthermore, we found that HDPs represent a gender-specific risk factor for future kidney dysfunction, underscoring the importance of incorporating obstetric history into the routine CKD assessments. Heightened awareness among healthcare providers regarding the link between HDPs and CKD can contribute significantly to the prevention of long-term kidney complications in women with a history of HDPs.

Notably, functional annotation of the genetic variants associated with HDPs revealed several genes involved in biological pathways crucial to both hypertensive pregnancy and renal physiology. Among these, *PLCE1* (phospholipase C epsilon 1) (rs10882398) is particularly significant as mutations in this gene are a well-established cause of early-onset nephrotic syndrome. This directly links it to

podocyte function and glomerular integrity, suggesting a compelling mechanism for albuminuria due to disruption of the renal filtration barrier [39,40]. Similarly, *TCF7L2* (transcription factor 7 like 2) (rs6585195), a major diabetes risk gene, indicates a potential shared genetic predisposition between HDPs and diabetic nephropathy, possibly mediated through metabolic disturbances influencing renal health [41]. Additional variants highlight more pathways: *PREX1* (phosphatidylinositol-3,4,5-trisphosphate-dependent Rac exchange factor 1) (rs2208589), involved in Angiotensin II signaling and endothelial function [42], offers a plausible link to the systemic endothelial dysfunction characteristic of preeclampsia. Likewise, *MTHFR* (methylenetetrahydrofolate reductase) (rs13306561), critical in folate metabolism and associated with hyperhomocysteinemia [43], may contribute to long-term vascular and renal injury via endothelial dysfunction and thrombosis. Furthermore, *THADA* (THADA armadillo repeat containing) (represented by multiple SNPs), associated with apoptosis regulation, diabetes, and polycystic ovary syndrome [44], suggests potential roles in both placental development and podocyte survival. Together, these genetic insights strengthen the biological plausibility of a causal pathway from HDPs to albuminuria, implicating podocyte injury, endothelial damage, and metabolic dysregulation as key mediators.

4.3 Limitations

Although our MR study addressed many limitations commonly found in observational designs, the results should be interpreted with caution due to constraints specific to our study and the broader MR framework. First, the use of GWAS data that included various forms of CKD may have introduced heterogeneity in the outcome, potentially affecting the robustness of our findings. Second, the lack of detailed information on preeclampsia severity and gestational age at onset prevented meaningful stratification of outcomes, which is essential for clinical translation. Future investigations into the relationship between HDPs and renal function should therefore incorporate more refined clinical phenotypes. Third, reliance on registry-based ICD codes for defining exposure may have resulted in underascertainment of subclinical or mild HDP cases and constrained our ability to evaluate dose-response relationships. Fourth, the number of strong genetic instruments varied across HDP subtypes, with fewer instruments available for gestational hypertension compared to preeclampsia. This likely reduced statistical power to detect significant associations for gestational hypertension, as suggested by the borderline association observed with albuminuria. As GWAS sample sizes for these phenotypes increase, future studies will be better equipped to examine the causal effects of specific HDP subtypes. Additionally, statistical power remained limited for certain analyses, such as between preeclampsia/eclampsia and eGFR, which may have hindered the de-

tection of a true causal effect despite a large outcome sample size (**Supplementary Table 7**). Finally, the use of mixed-sex outcome data, along with the absence of sex-stratified summary statistics, likely diluted effect estimates due to the inclusion of males who are not at risk for female-specific exposures. Eliminating this sex-related bias was not feasible.

5. Conclusions

To our knowledge, this study represents the first application of MR to examine the potential causal effect of HDPs on long-term renal outcomes. Incorporating a wide range of sensitivity analyses to validate the robustness of the results. This study found that the genetic proxy for HDPs was linked to maternal long-term albuminuria, and this association appeared to be independent of BMI, smoking, and T2D. Although the observed genetic association is modest in magnitude, it represents the lifelong, population-level effect of HDPs on albuminuria risk. Our findings provide novel evidence strengthening the causality between HDPs and long-term maternal kidney dysfunction.

Availability of Data and Materials

The GWAS summary statistics analyzed in this study are publicly available for download in the FinnGen (<https://www.finngen.fi/en>), CKDGen (<https://ckdgen.imbi.uni-freiburg.de/>), GIANT (https://giant-consortium.web.broadinstitute.org/GIANT_consortium), deCODE (<https://www.decode.com/>), and DIAMANTE (<https://diagram-consortium.org/>) consortia, as detailed in Table 1.

Author Contributions

XL and YX designed the study and supervised its implementation. HZ, PW, SS, QX, XZ, SC and ZW performed the data curation. HZ and YZ conducted the analysis. SS and HZ drafted the original version of the manuscript. HZ, YZ, QX, XZ and SC provided data interpretation, critical review, and commentary to the revised versions of the manuscript. All authors contributed to critical revision of the manuscript for important intellectual content. All authors read and approved the final manuscript. All authors have participated sufficiently in the work and agreed to be accountable for all aspects of the work.

Ethics Approval and Consent to Participate

Not applicable.

Acknowledgment

We extend our deep appreciation to the participants and researchers involved in the FinnGen study for their indispensable contributions to our research. We would like to acknowledge the UK Biobank (<https://www.ukbiobank.ac.uk/>) for providing the summary statistics data for our analysis. We also acknowledge and

appreciate the generosity of the GIANT, DIAMANTE, CKDGen, and deCODE consortia for sharing their invaluable resources.

Funding

This work was supported by Shenzhen Medical Research Fund (B2404004, A2303073), National Key Research and Development Program (2021YFC2701600), Shenzhen Science and Technology Program (RCYX20231211090407012, JCYJ20240813115007009, JCYJ20240813115015019, JCYJ20240813115014018), National Natural Science Foundation of China (82371698, 82502041), Medical Scientific Research Foundation of Guangdong Province (B2025763).

Conflict of Interest

The authors declare no conflict of interest.

Supplementary Material

Supplementary material associated with this article can be found, in the online version, at <https://doi.org/10.31083/CEOG44838>.

References

- [1] Say L, Chou D, Gemmill A, Tunçalp Ö, Moller AB, Daniels J, *et al.* Global causes of maternal death: a WHO systematic analysis. *The Lancet. Global Health*. 2014; 2: e323–e333. [https://doi.org/10.1016/S2214-109X\(14\)70227-X](https://doi.org/10.1016/S2214-109X(14)70227-X).
- [2] Chappell LC, Cluver CA, Kingdom J, Tong S. Pre-eclampsia. *Lancet* (London, England). 2021; 398: 341–354. [https://doi.org/10.1016/S0140-6736\(20\)32335-7](https://doi.org/10.1016/S0140-6736(20)32335-7).
- [3] Rana S, Lemoine E, Granger JP, Karumanchi SA. Preeclampsia: Pathophysiology, Challenges, and Perspectives. *Circulation Research*. 2019; 124: 1094–1112. <https://doi.org/10.1161/CIRCRESAHA.118.313276>.
- [4] Kattah A. Preeclampsia and Kidney Disease: Deciphering Cause and Effect. *Current Hypertension Reports*. 2020; 22: 91. <https://doi.org/10.1007/s11906-020-01099-1>.
- [5] Kristensen JH, Basit S, Wohlfahrt J, Damholt MB, Boyd HA. Pre-eclampsia and risk of later kidney disease: nationwide cohort study. *BMJ* (Clinical Research Ed.). 2019; 365: i1516. <https://doi.org/10.1136/bmj.i1516>.
- [6] Barrett PM, McCarthy FP, Kublickiene K, Cormican S, Judge C, Evans M, *et al.* Adverse Pregnancy Outcomes and Long-term Maternal Kidney Disease: A Systematic Review and Meta-analysis. *JAMA Network Open*. 2020; 3: e1920964. <https://doi.org/10.1001/jamanetworkopen.2019.20964>.
- [7] Srialuri N, Surapaneni A, Chang A, Mackeen AD, Paglia MJ, Grams ME. Preeclampsia and Long-term Kidney Outcomes: An Observational Cohort Study. *American Journal of Kidney Diseases: the Official Journal of the National Kidney Foundation*. 2023; 82: 698–705. <https://doi.org/10.1053/j.ajkd.2023.04.010>.
- [8] Shapiro J, Ray JG, McArthur E, Jeyakumar N, Chanchlani R, Harel Z, *et al.* Risk of Acute Kidney Injury After Hypertensive Disorders of Pregnancy: A Population-Based Cohort Study. *American Journal of Kidney Diseases: the Official Journal of the National Kidney Foundation*. 2022; 79: 561–569. <https://doi.org/10.1053/j.ajkd.2021.07.017>.
- [9] Kattah AG, Asad R, Scantlebury DC, Bailey KR, Wiste HJ, Hunt

SC, *et al.* Hypertension in pregnancy is a risk factor for microalbuminuria later in life. *Journal of Clinical Hypertension (Greenwich, Conn.)*. 2013; 15: 617–623. <https://doi.org/10.1111/jch.12116>.

[10] Vikse BE, Irgens LM, Leivestad T, Skjaerven R, Iversen BM. Preeclampsia and the risk of end-stage renal disease. *The New England Journal of Medicine*. 2008; 359: 800–809. <https://doi.org/10.1056/NEJMoa0706790>.

[11] Khashan AS, Evans M, Kubickas M, McCarthy FP, Kenny LC, Stenvinkel P, *et al.* Preeclampsia and risk of end stage kidney disease: A Swedish nationwide cohort study. *PLoS Medicine*. 2019; 16: e1002875. <https://doi.org/10.1371/journal.pmed.1002875>.

[12] Sanderson E, Glymour MM, Holmes MV, Kang H, Morrison J, Munafò MR, *et al.* Mendelian randomization. *Nature Reviews. Methods Primers*. 2022; 2: 6. <https://doi.org/10.1038/s43586-021-00092-5>.

[13] Kurki MI, Karjalainen J, Palta P, Sipilä TP, Kristiansson K, Donner KM, *et al.* FinnGen provides genetic insights from a well-phenotyped isolated population. *Nature*. 2023; 613: 508–518. <https://doi.org/10.1038/s41586-022-05473-8>.

[14] Wuttke M, Li Y, Li M, Sieber KB, Feitosa MF, Gorski M, *et al.* A catalog of genetic loci associated with kidney function from analyses of a million individuals. *Nature Genetics*. 2019; 51: 957–972. <https://doi.org/10.1038/s41588-019-0407-x>.

[15] Haas ME, Aragam KG, Emdin CA, Bick AG, International Consortium for Blood Pressure, Hemani G, *et al.* Genetic Association of Albuminuria with Cardiometabolic Disease and Blood Pressure. *American Journal of Human Genetics*. 2018; 103: 461–473. <https://doi.org/10.1016/j.ajhg.2018.08.004>.

[16] Pulit SL, Stoneman C, Morris AP, Wood AR, Glastonbury CA, Tyrrell J, *et al.* Meta-analysis of genome-wide association studies for body fat distribution in 694 649 individuals of European ancestry. *Human Molecular Genetics*. 2019; 28: 166–174. <https://doi.org/10.1093/hmg/ddy327>.

[17] Mahajan A, Spracklen CN, Zhang W, Ng MCY, Petty LE, Kitajima H, *et al.* Multi-ancestry genetic study of type 2 diabetes highlights the power of diverse populations for discovery and translation. *Nature Genetics*. 2022; 54: 560–572. <https://doi.org/10.1038/s41588-022-01058-3>.

[18] Liu M, Jiang Y, Wedow R, Li Y, Brazel DM, Chen F, *et al.* Association studies of up to 1.2 million individuals yield new insights into the genetic etiology of tobacco and alcohol use. *Nature Genetics*. 2019; 51: 237–244. <https://doi.org/10.1038/s41588-018-0307-5>.

[19] Levey AS, Stevens LA, Schmid CH, Zhang YL, Castro AF, 3rd, Feldman HI, *et al.* A new equation to estimate glomerular filtration rate. *Annals of Internal Medicine*. 2009; 150: 604–612. <https://doi.org/10.7326/0003-4819-150-9-200905050-00006>.

[20] Schwartz GJ, Schneider MF, Maier PS, Moxey-Mims M, Dharnidharka VR, Warady BA, *et al.* Improved equations estimating GFR in children with chronic kidney disease using an immunonephelometric determination of cystatin C. *Kidney International*. 2012; 82: 445–453. <https://doi.org/10.1038/ki.2012.169>.

[21] Wikström AK, Stephansson O, Cnattingius S. Tobacco use during pregnancy and preeclampsia risk: effects of cigarette smoking and snuff. *Hypertension (Dallas, Tex.: 1979)*. 2010; 55: 1254–1259. <https://doi.org/10.1161/HYPERTENSIONAHA.109.147082>.

[22] Magee LA, Nicolaides KH, von Dadelszen P. Preeclampsia. *The New England Journal of Medicine*. 2022; 386: 1817–1832. <https://doi.org/10.1056/NEJMra2109523>.

[23] Boyko EJ, Seelig AD, Jacobson IG, Hooper TI, Smith B, Smith TC, *et al.* Sleep characteristics, mental health, and diabetes risk: a prospective study of U.S. military service members in the Millennium Cohort Study. *Diabetes Care*. 2013; 36: 3154–3161. <https://doi.org/10.2337/dc13-0042>.

[24] Lawlor DA, Benfield L, Logue J, Tilling K, Howe LD, Fraser A, *et al.* Association between general and central adiposity in childhood, and change in these, with cardiovascular risk factors in adolescence: prospective cohort study. *BMJ (Clinical Research Ed.)*. 2010; 341: c6224. <https://doi.org/10.1136/bmj.c6224>.

[25] Frayling TM, Timpson NJ, Weedon MN, Zeggini E, Freathy RM, Lindgren CM, *et al.* A common variant in the FTO gene is associated with body mass index and predisposes to childhood and adult obesity. *Science (New York, N.Y.)*. 2007; 316: 889–894. <https://doi.org/10.1126/science.1141634>.

[26] Hemani G, Zheng J, Elsworth B, Wade KH, Haberland V, Baird D, *et al.* The MR-Base platform supports systematic causal inference across the human genome. *eLife*. 2018; 7: e34408. <https://doi.org/10.7554/eLife.34408>.

[27] Hartwig FP, Davey Smith G, Bowden J. Robust inference in summary data Mendelian randomization via the zero modal pleiotropy assumption. *International Journal of Epidemiology*. 2017; 46: 1985–1998. <https://doi.org/10.1093/ije/dyx102>.

[28] Bowden J, Davey Smith G, Burgess S. Mendelian randomization with invalid instruments: effect estimation and bias detection through Egger regression. *International Journal of Epidemiology*. 2015; 44: 512–525. <https://doi.org/10.1093/ije/dyv080>.

[29] Verbanck M, Chen CY, Neale B, Do R. Detection of widespread horizontal pleiotropy in causal relationships inferred from Mendelian randomization between complex traits and diseases. *Nature Genetics*. 2018; 50: 693–698. <https://doi.org/10.1038/s41588-018-0099-7>.

[30] Greco M FD, Minelli C, Sheehan NA, Thompson JR. Detecting pleiotropy in Mendelian randomisation studies with summary data and a continuous outcome. *Statistics in Medicine*. 2015; 34: 2926–2940. <https://doi.org/10.1002/sim.6522>.

[31] McDonald SD, Han Z, Walsh MW, Gerstein HC, Devereaux PJ. Kidney disease after preeclampsia: a systematic review and meta-analysis. *American Journal of Kidney Diseases: the Official Journal of the National Kidney Foundation*. 2010; 55: 1026–1039. <https://doi.org/10.1053/j.ajkd.2009.12.036>.

[32] Garovic VD, White WM, Vaughan L, Saiki M, Parashuram S, Garcia-Valencia O, *et al.* Incidence and Long-Term Outcomes of Hypertensive Disorders of Pregnancy. *Journal of the American College of Cardiology*. 2020; 75: 2323–2334. <https://doi.org/10.1016/j.jacc.2020.03.028>.

[33] Arechvo A, Wright A, Syngelaki A, von Dadelszen P, Magee LA, Akolekar R, *et al.* Incidence of pre-eclampsia: effect of deprivation. *Ultrasound in Obstetrics & Gynecology: the Official Journal of the International Society of Ultrasound in Obstetrics and Gynecology*. 2023; 61: 26–32. <https://doi.org/10.1002/uog.26084>.

[34] Arechvo A, Voicu D, Gil MM, Syngelaki A, Akolekar R, Nicolaides KH. Maternal race and pre-eclampsia: Cohort study and systematic review with meta-analysis. *BJOG: an International Journal of Obstetrics and Gynaecology*. 2022; 129: 2082–2093. <https://doi.org/10.1111/1471-0528.17240>.

[35] Piccoli GB, Attini R, Vasario E, Conijn A, Biolcati M, D'Amico F, *et al.* Pregnancy and chronic kidney disease: a challenge in all CKD stages. *Clinical Journal of the American Society of Nephrology: CJASN*. 2010; 5: 844–855. <https://doi.org/10.2215/CJN.07911109>.

[36] Tangren JS, Powe CE, Ankers E, Ecker J, Bramham K, Hladunewich MA, *et al.* Pregnancy Outcomes after Clinical Recovery from AKI. *Journal of the American Society of Nephrology: JASN*. 2017; 28: 1566–1574. <https://doi.org/10.1681/ASN.2016070806>.

[37] Fisher KA, Luger A, Spargo BH, Lindheimer MD. Hypertension in pregnancy: clinical-pathological correlations and remote

prognosis. *Medicine*. 1981; 60: 267–276.

[38] Garovic VD, Wagner SJ, Petrovic LM, Gray CE, Hall P, Sugimoto H, *et al.* Glomerular expression of nephrin and synaptopodin, but not podocin, is decreased in kidney sections from women with preeclampsia. *Nephrology, Dialysis, Transplantation: Official Publication of the European Dialysis and Transplant Association - European Renal Association*. 2007; 22: 1136–1143. <https://doi.org/10.1093/ndt/gfl711>.

[39] Atchison DK, O'Connor CL, Menon R, Otto EA, Ganesh SK, Wiggins RC, *et al.* Hypertension induces glomerulosclerosis in phospholipase C-ε1 deficiency. *American Journal of Physiology. Renal Physiology*. 2020; 318: F1177–F1187. <https://doi.org/10.1152/ajprenal.00541.2019>.

[40] Hinkes B, Wiggins RC, Gbadegesin R, Vlangos CN, Seelow D, Nürnberg G, *et al.* Positional cloning uncovers mutations in PLCE1 responsible for a nephrotic syndrome variant that may be reversible. *Nature Genetics*. 2006; 38: 1397–1405. <https://doi.org/10.1038/ng1918>.

[41] Grant SFA. The *TCF7L2* Locus: A Genetic Window Into the Pathogenesis of Type 1 and Type 2 Diabetes. *Diabetes Care*. 2019; 42: 1624–1629. <https://doi.org/10.2337/dc19-0001>.

[42] Naikawadi RP, Cheng N, Vogel SM, Qian F, Wu D, Malik AB, *et al.* A critical role for phosphatidylinositol (3,4,5)-trisphosphate-dependent Rac exchanger 1 in endothelial junction disruption and vascular hyperpermeability. *Circulation Research*. 2012; 111: 1517–1527. <https://doi.org/10.1161/CIRCRESAHA.112.273078>.

[43] Blom HJ, Verhoef DP. Hyperhomocysteinemia, MTHFR, and risk of vascular disease. *Circulation*. 2000; 101: E171; author reply E173. <https://doi.org/10.1161/01.cir.101.16.e171>.

[44] Tian Y, Li J, Su S, Cao Y, Wang Z, Zhao S, *et al.* PCOS-GWAS Susceptibility Variants in *THADA*, *INSR*, *TOX3*, and *DENND1A* Are Associated With Metabolic Syndrome or Insulin Resistance in Women With PCOS. *Frontiers in Endocrinology*. 2020; 11: 274. <https://doi.org/10.3389/fendo.2020.00274>.